

## Challenging Volatility Myths: Structural and Temporal Dynamics in MENA's Growth-Openness Nexus

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### ABSTRACT

This study addresses two critical gaps in the development literature: whether macroeconomic volatility necessarily undermines long-term growth or can generate productivity-enhancing dynamics under specific conditions, and whether trade openness amplifies or mitigates the volatility–growth nexus across structurally heterogeneous economies. By systematically comparing oil exporters and importers in the MENA region during 1990–2023, the research provides the first comprehensive regional analysis of how resource endowment fundamentally conditions the three-way interaction between volatility, openness, and growth in this strategically important yet understudied region. The rapid expansion of global trade has created a paradoxical dynamic: while openness provides growth opportunities through market access and knowledge spillovers, it also heightens exposure to external shocks and global volatility. Theoretical predictions range from neoclassical efficiency gains to Schumpeterian creative destruction mechanisms, with empirical evidence remaining inconclusive across different contexts. Recent disruptions such as the COVID-19 pandemic and geopolitical conflicts have revived debates on whether openness amplifies fragility or fosters resilience. To investigate these issues, a comprehensive unbalanced panel of 18 MENA countries covering 612 country-year observations is constructed, and System-GMM estimation is applied to address endogeneity concerns. A novel policy-adjusted Trade Openness Policy (TOP) indicator, derived from gravity model decomposition, isolates discretionary policy effects from structural determinants of trade flows. The indicator demonstrates high explanatory power ( $R^2 > 0.7$ ) while maintaining weak correlation with conventional trade ratios ( $\rho < 0.3$ ), ensuring cleaner identification of policy effects. Sequential specifications progress from baseline volatility–growth estimations to interactive moderation effects, with the Johnson–Neyman technique identifying precise threshold values where openness significantly alters volatility–growth relationships. Comprehensive robustness checks account for temporal structural breaks, alternative volatility measures, and conflict-related disruptions. Results fundamentally challenge conventional wisdom by demonstrating a positive volatility–growth relationship ( $\beta = 0.0526$ ,  $p < 0.05$ ), consistent with Schumpeterian “creative destruction” dynamics in resource-dependent developing contexts. Trade openness displays weak direct growth effects but proves decisive in moderation, with critical temporal variation: openness provided stabilizing effects before 2008 ( $\beta = 0.0104$ ,  $p < 0.01$ ) but became destabilizing after 2008 ( $\beta = -0.0109$ ,  $p < 0.001$ ), reflecting fundamental changes in global economic architecture. Government effectiveness emerges as the most robust growth determinant across all specifications, with one standard deviation improvements translating into 40–46% higher per capita income over two decades. The study makes three distinct contributions: it provides the first systematic region-wide analysis of volatility–openness–growth interactions in MENA, reveals pronounced structural heterogeneity between exporters and importers, introduces methodological innovation through policy-based openness measurement and threshold identification techniques that offer replicable tools for future research, and delivers policy-relevant

insights demonstrating that optimal development strategies differ fundamentally between resource-rich exporters and diversified importers, with timing and institutional context proving critical for successful integration.

**Keywords:** Macroeconomic volatility; Trade openness; Economic growth; MENA economies; System-GMM; Johnson-Neyman technique

## INTRODUCTION

The unprecedented expansion of global trade, consistently outpacing world GDP growth in recent decades, has profoundly reshaped how economic shocks propagate across borders (Baldwin & Lopez-Gonzalez, 2015). Greater integration has created a paradoxical dynamic: while openness provides access to markets, knowledge spillovers, and growth-enhancing specialization, it simultaneously heightens exposure to external shocks and global volatility (Kose et al., 2003; di Giovanni & Levchenko, 2009).

Theoretical predictions about this trade-off remain contested. Neoclassical models emphasize efficiency gains but highlight increased vulnerability from sectoral specialization (Helpman & Krugman, 1985). Endogenous growth frameworks stress that technology diffusion and knowledge transfers may offset volatility costs (Grossman & Helpman, 1991; Rivera-Batiz & Romer, 1991). A third perspective suggests that moderate volatility itself can stimulate long-run growth by fostering creative destruction and resource reallocation (Aghion & Saint-Paul, 1998; Imbs, 2007). Empirical findings are equally inconclusive, with evidence for both stabilizing effects of openness through diversification (Bekaert et al., 2006) and destabilizing effects via synchronization with global cycles (Razin & Rose, 1994; Calderon et al., 2007).

Recent global events have reopened this debate. The COVID-19 pandemic exposed critical vulnerabilities in global supply chains, while the war in Ukraine underscored the geopolitical risks of external dependence (Antràs, 2021). These shocks highlight the need to reassess whether openness amplifies fragility or provides buffers against cyclical instability.

Against this backdrop, our study addresses two unresolved questions. First, does macroeconomic volatility necessarily undermine long-term growth, or can certain structural contexts—even volatility itself—generate productivity-enhancing dynamics? Second, does trade openness moderate the volatility–growth relationship, and does this role differ across structurally heterogeneous economies? Unlike much of the existing literature, which implicitly assumes volatility is detrimental, we explicitly test whether instability may, under specific conditions, coincide with or even promote growth.

We focus on the Middle East and North Africa (MENA) region during 1990–2023, a unique testing ground that combines resource-rich oil exporters—highly dependent on global commodity cycles—with more diversified oil importers facing different structural vulnerabilities (Nugent & Pesaran, 2007; Arezki & Nabli, 2012). The region is particularly illustrative given its recurrent exposure to oil price shocks, political crises, and global downturns.

This paper makes three contributions.

1. Revisiting the volatility–growth link: We assess whether cyclical instability systematically hampers growth or, consistent with Schumpeterian theory, may foster reallocation dynamics that support long-run development. Our findings confirm that in the MENA region, moderate volatility can be positively correlated with growth, particularly among oil exporters.
2. Reassessing the role of trade openness: We move beyond linear assumptions by testing whether openness moderates the volatility–growth nexus, using the Johnson–Neyman technique to detect thresholds where openness shifts from amplifying to stabilizing. While the direct effect of openness on growth appears weak, its moderating role proves more consequential, with heterogeneous effects across exporters and importers.
3. Accounting for MENA heterogeneity: By distinguishing oil exporters from importers, we provide the first systematic evidence on how resource endowment and structural characteristics condition the growth–volatility–openness nexus. Our results reveal that trade openness amplifies stability gains for importers, while its role among exporters is more limited.

Together, these contributions move beyond conventional expectations of openness-driven growth and volatility-driven fragility. They align the empirical analysis with the realities of the MENA region, where evidence suggests nonlinear and context-dependent interactions: volatility may coincide with stronger growth, and the benefits of openness are neither automatic nor uniformly stabilizing.

The remainder of the paper is structured as follows. Section 2 reviews the theoretical and empirical literature and develops testable hypotheses. Section 3 introduces the dataset, descriptive statistics, and stylized facts. Section 4 outlines the econometric methodology and identification strategy. Section 5 reports empirical results, robustness checks, and threshold analyses. Section 6 concludes with policy implications and future research directions.

## THEORETICAL CONSIDERATIONS AND LITERATURE REVIEW

The nexus between macroeconomic volatility, economic growth, and trade openness remains one of the most theoretically rich yet empirically contested topics in modern macroeconomics. This relationship carries profound implications for understanding growth dynamics and for designing policy frameworks in an increasingly integrated global economy. Despite decades of theoretical development and empirical testing, fundamental questions remain unresolved: Does macroeconomic instability systematically hinder long-run growth, or can it under certain conditions foster productivity-enhancing structural transformation? Does trade openness amplify vulnerability to external shocks, or does it strengthen resilience through diversification and competition?

This section synthesizes theoretical foundations and empirical evidence across three research streams that directly inform our investigation. We first revisit the evolving understanding of the growth–volatility relationship, moving beyond traditional dichotomies. We then examine how trade openness affects economic performance, highlighting the conditioning role of institutions and structural characteristics. Finally, we assess how openness moderates the volatility–growth nexus, with emphasis on the mechanisms that determine whether integration stabilizes or destabilizes.

In the Real Business Cycle framework (Kydland & Prescott, 1982), volatility was treated as transitory fluctuations around an exogenous growth trend, with no permanent impact. Later work identified several mechanisms through which volatility can affect long-term growth. Investment uncertainty discourages capital accumulation (Bernanke, 1983; Dixit & Pindyck, 1994). Financial accelerator effects amplify downturns via credit constraints, with hysteresis consequences for productivity (Bernanke et al., 1999; Aghion et al., 2010). Human capital channels show recessions reduce incentives and resources for education and training, lowering future productivity (Acemoglu & Scott, 1997; Krebs, 2003). Conversely, alternative theories suggest volatility may promote growth. Creative destruction reallocates resources to more productive firms (Caballero & Hammour, 1994; Aghion & Saint-Paul, 1998), while cleansing effects purge inefficient firms, improving aggregate efficiency (Barlevy, 2002; Ouyang, 2009). Whether volatility is harmful or beneficial depends on an economy's absorptive capacity—institutional quality, financial depth, and labor market flexibility (Aghion et al., 2018).

Empirical evidence reflects this theoretical tension. Early studies (Kormendi & Meguire, 1985; Grier & Tullock, 1989) found positive correlations between volatility and growth, while Ramey & Ramey (1995) documented a robust negative relationship, especially in developing economies. Subsequent work identified conditioning factors: weak financial systems exacerbate costs (Hnatkovska & Loayza, 2005; Kose et al., 2006), threshold effects matter (Baum et al., 2013), volatility in tradables is more damaging (Loayza & Raddatz, 2007), and strong governance mitigates risks. Overall, while the dominant narrative emphasizes volatility's costs, evidence supports the possibility that instability may, under specific conditions, coincide with higher growth—an argument directly tested in this study. Consistent with this ambiguity, our empirical evidence for MENA confirms that volatility can be positively associated with growth, particularly among oil exporters, suggesting that moderate instability may carry expansionary rather than purely harmful effects.

Turning to the second pillar of our theoretical framework, the trade openness–growth literature reveals equally complex patterns. Openness may stimulate growth via static efficiency (Krugman, 1980), technology diffusion (Coe & Helpman, 1995), learning-by-exporting (Clerides et al., 1998), and competition-driven productivity (Melitz, 2003). Financial integration enhances capital access and risk-sharing (Bekaert et al., 2005), while institutional upgrading follows alignment with global standards (Do & Levchenko, 2007). Yet openness also exposes economies to capital flow volatility, contagion, and sectoral specialization risks. Early studies (Edwards, 1998; Frankel & Romer, 1999) reported robust pro-growth effects, but critiques (Rodriguez & Rodrik, 2001) questioned measurement and endogeneity. More recent evidence highlights conditionality: institutional strength, diversification (Balavac & Pugh, 2016), and thresholds where excessive openness yields diminishing or negative returns. Our findings align with this nuanced perspective, showing that the direct growth effects of openness are weak once structural and institutional heterogeneity are accounted for. Indeed, openness coefficients are negative for exporters and only weakly positive for importers, underscoring the limits of a universal "openness dividend" in the MENA region.

The third dimension of our analysis concerns how openness moderates the volatility–growth relationship, where competing theoretical perspectives emerge. The vulnerability hypothesis posits that openness amplifies volatility's costs through terms-of-trade shocks, sectoral specialization, and reduced policy autonomy (Rodrik, 1998; Easterly et al., 2001). The diversification hypothesis suggests the opposite—that openness stabilizes growth via risk-sharing and market broadening (Kose et al., 2003). The outcome depends on institutional and financial capacity (Caselli et al., 2015). Empirical evidence is divided: openness amplifies volatility in poorly diversified economies with weak institutions (Kose et al., 2006), while it mitigates volatility where governance and financial depth are strong (Aghion et al., 2018). Sectoral structure is decisive (Loayza & Raddatz, 2007). In MENA, structural

heterogeneity implies contrasting effects: oil exporters, concentrated in volatile commodities with weaker institutions, may experience amplified vulnerability, while diversified importers may benefit from stabilization but remain exposed to energy costs and financing shocks. Our results reinforce this contingency: overall moderation effects are weak, but importers display a significant positive interaction between openness and volatility, suggesting that trade liberalization can provide stabilizing benefits when institutional and structural conditions are supportive.

Despite abundant scholarship, three research gaps persist that our study addresses. First, MENA-specific analyses remain scarce, despite the region's recurrent shocks and structural heterogeneity. Second, evidence on non-linear moderation is limited, with few studies applying threshold techniques to identify precise ranges where openness effects change. Third, conventional openness measures conflate structural trade determinants with discretionary policy choices, obscuring the true policy impact. Our study fills these gaps through several methodological innovations. We provide the first systematic MENA analysis distinguishing oil exporters from importers, recognizing their fundamentally different economic structures. We apply System-GMM estimation to handle endogeneity concerns while employing Johnson–Neyman threshold techniques to identify precise moderation ranges where openness effects shift. Finally, we construct a policy-adjusted openness indicator that isolates discretionary policy effects from structural trade determinants, providing clearer guidance for policy design.

## Data and Stylized Facts

Our empirical analysis employs a comprehensive unbalanced panel dataset encompassing 18 MENA countries over 1990–2023, yielding 612 country-year observations. The temporal scope strategically captures major macroeconomic disruptions—the Gulf War (1990–1991), Asian Financial Crisis (1997–1998), Global Financial Crisis (2008–2009), Arab Spring (2010–2012), and COVID-19 pandemic (2020–2021)—providing substantial variation in growth performance and volatility sources crucial for robust econometric identification. Volatility is measured as the standard deviation of annual GDP per capita growth rates over five-year rolling windows, following established practice (Ramey & Ramey, 1995). For MENA countries, volatility is sourced from the pre-calculated instability metric in the dataset, while for OECD countries, it is computed directly from annual growth rates.

The sample is partitioned into two economically distinct sub-groups reflecting the region's fundamental structural heterogeneity. MENA Oil Exporters (10 countries: Algeria, Bahrain, Iraq, Iran, Kuwait, Libya, Oman, Qatar, Saudi Arabia, UAE) are characterized by substantial hydrocarbon reserves (>10% of global totals), fiscal dependence on energy revenues (>60% of government income), export concentration in oil and gas (>70% of merchandise exports), and pronounced exposure to commodity price volatility with limited economic diversification. MENA Oil Importers (8 countries: Egypt, Jordan, Lebanon, Mauritania, Morocco, Syria, Tunisia, Turkey) exhibit negligible hydrocarbon endowments, diversified production structures spanning manufacturing and services, greater reliance on manufactured exports and tourism revenues, and vulnerability to energy price fluctuations primarily through import cost channels.

This classification follows established IMF and World Bank taxonomies, with countries classified as oil exporters if net energy exports exceed 10% of total exports and hydrocarbon revenues constitute >20% of government income over the sample period. This threshold-based approach ensures consistent classification while acknowledging dynamic transitions due to production changes and political disruptions.

Data combines multiple authoritative sources to maximize coverage and reliability. Real GDP per capita (constant 2015 US\$) derives from World Bank World Development Indicators. Trade flows are sourced from IMF Direction of Trade Statistics and UN Comtrade. Institutional quality indicators come from World Bank Worldwide Governance Indicators and Polity IV Project. Financial development metrics are drawn from the Global Financial Development Database, while oil prices originate from IMF Primary Commodity Prices database. The MENA dataset includes additional governance indicators (control of corruption, government effectiveness, rule of law, regulatory quality, political stability, voice and accountability), human capital, terms of trade, and a policy-adjusted trade openness indicator (TOP).

Table 1 presents comprehensive descriptive statistics comparing MENA sub-groups with 34 OECD economies as a benchmark. The volatility-to-growth ratio provides a novel efficiency metric—lower ratios indicate countries achieve growth with proportionally less macroeconomic instability, capturing the trade-off between expansion and stability that lies at the heart of our theoretical framework. The descriptive analysis reveals fundamental patterns challenging conventional macroeconomic wisdom. MENA economies slightly underperform OECD countries in average growth rates (1.25% versus 1.69%), reflecting catch-up dynamics, favorable demographics, and resource-driven expansion characteristic of middle-income economies. However, this growth entails substantially higher instability, with MENA volatility averaging 6.64% compared to 3.25% for OECD economies. The volatility-to-growth ratio for MENA (8.87) significantly exceeds OECD levels (2.12), indicating that MENA growth systematically accompanies disproportionate instability—an "efficiency gap" suggesting insufficient institutional and structural capabilities for stable growth (Aghion et al., 2018).

Extreme cases illuminate these dynamics. Kuwait, Libya, and Iraq exhibit volatility-to-growth ratios exceeding 3.0 (9.70, 23.12, and 6.55, respectively), reflecting geopolitical conflicts, institutional fragility, and excessive commodity dependence. Libya's extreme volatility (21.89%) and Iraq's (18.99%) stem from prolonged conflicts and oil price shocks, while Kuwait's (12.59%) reflects Gulf War disruptions and commodity dependence, demonstrating how resource abundance becomes instability rather than prosperity—consistent with resource curse literature (Sachs & Warner, 2001; Mehlum et al., 2006).

The oil exporter–importer comparison reveals systematic differences reinforced by our econometric results. While exporters achieve lower average growth (0.96% versus 1.44%), they exhibit dramatically higher volatility (8.19% versus 4.18%) and volatility-to-growth ratios (14.67 versus 1.61). Our regression findings confirm this duality: the growth–volatility relationship is positive among exporters, reflecting persistent boom–bust cycles driven by resource dependence. Among importers, however, the relationship appears weak and statistically insignificant, suggesting that diversified structures provide resilience but without generating a clear stabilizing effect. This divergence underscores how resource dependence alters growth processes through terms-of-trade volatility, Dutch disease effects, and procyclical fiscal policies, while diversified economies partially decouple growth from volatility.

Figure 1 presents scatter plots examining empirical growth–volatility relationships across country groupings, providing compelling preliminary evidence that structural characteristics fundamentally shape this nexus in ways standard theoretical models fail to predict. The pooled MENA–OECD sample exhibits a weak positive correlation (0.079,  $p=0.576$ ), suggesting no strong evidence that higher growth requires volatility tolerance. The OECD subsample displays a positive growth–volatility correlation (0.241,  $p=0.169$ ), diverging from established theory predicting a negative relationship due to developed institutions and financial markets enabling stable growth (Ramey & Ramey, 1995; Hnatkovska & Loayza, 2005), though this result is not statistically significant. The full MENA sample shows a moderate positive growth–volatility correlation (0.333,  $p=0.177$ ), challenging conventional wisdom. Crucially, our empirical models highlight dual patterns: oil exporters exhibit a strong positive correlation (0.679,  $p=0.093$ ), consistent with regression evidence showing volatility associates with growth in resource-rich economies, though the relationship is marginally significant. Countries with the highest volatility (Kuwait, Libya, Iraq) record high growth during favorable periods, illustrating how resource windfalls fuel both instability and expansion. Conversely, oil importers display a moderate positive correlation (0.569,  $p=0.317$ ), indicating that structural diversification tempers volatility's adverse effects without fully decoupling growth from instability.

Preliminary analysis reveals correlations between a policy-adjusted trade openness measure (TOP) and economic outcomes across MENA sub-groups, nuancing the growth–openness–volatility nexus. For oil exporters, openness is negatively correlated with growth ( $-0.42$ ,  $p=0.225$ ), reflecting the dominance of resource trade and distortions from hydrocarbon dependence. For oil importers, openness also shows a negative correlation with growth ( $-0.37$ ,  $p=0.365$ ), suggesting that trade openness does not strongly drive growth in diversified economies. The openness–volatility relationship further highlights structural differences: for exporters, openness has a weak negative correlation with volatility ( $-0.11$ ,  $p=0.757$ ), indicating that higher openness does not significantly increase instability. For importers, the correlation is stronger and negative ( $-0.54$ ,  $p=0.165$ ), suggesting that diversified trade structures may reduce volatility, though the relationship remains statistically insignificant.

These stylized facts provide crucial guidance for our empirical methodology and confirm the need for heterogeneity-aware modeling. The pronounced exporter–importer divergence justifies separate estimation approaches rather than pooled analysis with dummy variables. Moreover, the presence of nonlinear and conditional patterns validates our application of threshold techniques and our focus on a policy-adjusted openness indicator. The MENA experience demonstrates how resource dependence and diversification shape the volatility–growth–openness nexus, underscoring the importance of context-specific analysis rather than universal generalizations from pooled cross-country studies.

## METHODOLOGY AND MODEL CONSTRUCTION

This section presents our econometric framework designed to rigorously test the three hypotheses derived from theoretical analysis while addressing key empirical challenges: endogeneity, parameter heterogeneity, and nonlinear threshold effects. Our sequential modeling approach builds complexity progressively, explicitly capturing the structural duality between oil exporters and importers observed in our descriptive evidence and confirmed by the regression results.

To evaluate whether macroeconomic instability systematically undermines economic growth (H1), we begin with a baseline specification:

$$y_{i,t} = \alpha_0 + \alpha_1 \text{Ins}_{i,t-1} + \rho' X_{i,t-1} + \beta_i + \beta_t + \mu_{i,t}$$

where  $y_{i,t}$  denotes real GDP per capita growth,  $\beta_t$  and  $\beta_i$  capture country and time fixed effects, and  $X_{i,t-1}$  represents lagged controls established in growth literature.

Following Ramey and Ramey (1995) and Hnatkovska and Loayza (2005), we employ two-step volatility construction to measure instability. First, we perform trend-cycle decomposition:

$$Y_{i,t} = \alpha_i + \beta_{it} + \gamma_i Y_{i,t-1} + \delta_i Y_{i,t-2} + \xi_{i,t}$$

Second, rolling volatility over  $k=5$  years:

$$Ins_{i,t} = \frac{1}{k} \sum_{s=0}^{k-1} \xi_{i,t-s}^2$$

This approach isolates genuine business cycle fluctuations from measurement error and structural breaks while allowing time-varying volatility.

Our control variable vector  $X_{i,t-1}$  includes four categories: (i) structural factors (initial income, population growth, investment rate), (ii) policy volatility (monetary and fiscal instability), (iii) external shocks (terms of trade, oil price volatility), and (iv) institutional quality (government effectiveness). One-period lags address simultaneity while reflecting realistic policy transmission.

To test whether trade openness independently enhances growth (H2), we extend our framework to:

$$y_{i,t} = \alpha_0 + \alpha_1 TOP_{i,t-1} + \rho' X_{i,t-1} + \beta_i + \beta_t + \mu_{i,t}$$

A key methodological innovation concerns our trade openness measurement. Conventional trade-to-GDP ratios conflate policy choices with structural factors. Following Pritchett (1996), we employ gravity-based decomposition. In the first stage, we predict structural trade exposure:

$$\log\left(\frac{Trade_{i,t}}{GDP_{i,t}}\right) = \beta_0 + \beta_1 \log(GDP_{i,t}) + \beta_2 \log(Pop_{i,t}) + \beta_3 Resource_i + \beta_4 \log(Distance_i) \\ + \beta_5 Landlocked_i + \beta_6 Colonial_i + \varepsilon_{i,t}$$

In the second stage, we construct our Trade Openness Policy (TOP) indicator:

$$TOP_{i,t} = \ln\left(\frac{Trade_{i,t}}{GDP_{i,t}}\right) - \ln\left(\widehat{\frac{Trade_{i,t}}{GDP_{i,t}}}\right)$$

Positive TOP values indicate more liberal policies than structural characteristics predict, while negative values suggest policy barriers.

Given the unique characteristics of MENA economies, we implement several region-specific adjustments. For oil exporters, we separate hydrocarbon and non-hydrocarbon trade to distinguish resource-driven openness from broader integration. Regional integration dummies capture Arab Monetary Fund and GCC effects. For conflict-affected countries (Iraq, Libya, Syria), we employ interpolation to avoid conflating war disruptions with policy choices.

To examine whether trade openness moderates the volatility-growth relationship (H3), we estimate our core interaction model:

$$y_{i,t} = \alpha_0 + \alpha_1 TOP_{i,t-1} + \alpha_2 Ins_{i,t-1} + \alpha_3 (TOP_{i,t-1} \times Ins_{i,t-1}) + \rho' X_{i,t-1} + \beta_i + \beta_t + \mu_{i,t}$$

The interaction coefficient  $\alpha_3$  distinguishes between competing theoretical hypotheses. The Diversification Hypothesis predicts  $\alpha_3 < 0$ , indicating that openness mitigates volatility's negative growth effects. The Vulnerability Hypothesis predicts  $\alpha_3 > 0$ , suggesting that openness amplifies volatility's adverse impact.

Recognizing that linear interaction models may inadequately capture complex relationships, we implement Johnson-Neyman analysis to identify threshold values where volatility effects become statistically significant. The conditional effect of volatility at different openness levels is:

$$\theta(TOP) = \alpha_2 + \alpha_3 \times TOP$$

Critical values solve:

$$\frac{|\theta(TOP)|}{SE[\theta(TOP)]} = 1.966$$

where

$$SE[\theta(TOP)] = \sqrt{\text{Var}(\alpha_2) + TOP^2 \times \text{Var}(\alpha_3) + 2 \times TOP \times \text{Cov}(\alpha_2, \alpha_3)}$$

This technique identifies optimal openness ranges for different economic structures and provides precise policy guidance.

Our estimation strategy proceeds through three sequential stages to ensure robustness. In the first stage, we employ baseline panel methods including Fixed Effects estimation controlling for time-invariant heterogeneity, Hausman tests for specification selection, and diagnostic testing for heteroskedasticity, serial correlation, and cross-sectional dependence. The second stage employs dynamic panel methods through System GMM estimation, which addresses endogeneity through internal instruments using deeper lags ( $t-3$  and beyond), external instruments (geographic variables, colonial ties), Hansen J-tests for instrument validity, and Arellano-Bond tests for serial correlation. The third stage conducts comprehensive robustness analysis using alternative volatility windows (3, 5, 7 years), different openness measures (export orientation, import penetration), sub-period analysis (pre/post-2008, pre/post-Arab Spring), and outlier exclusion with bootstrap standard errors.

Endogeneity concerns arise from multiple sources that our identification strategy must address. Reverse causation may occur as growth affects volatility through investment dynamics. Joint determination problems emerge when trade policy and growth strategies are simultaneously chosen. Omitted variables bias may result from unobserved institutional factors affecting both openness and growth. Our identification strategy tackles these challenges through System GMM, which exploits both within-country variation and cross-sectional differences. We employ deep lags using  $t-3$  and beyond to minimize contamination, geographic instruments including distance to major trading partners and landlocked status, and historical instruments such as colonial ties and neighboring countries' policies.

Given fundamental differences between oil exporters and importers revealed in Section 3, we estimate all models separately for each sub-group. This approach reflects empirical evidence showing that exporters display instability with positive and significant associations with growth reflecting boom-bust dynamics, but openness effects are negative and moderation is weak or insignificant. Importers show weak or absent instability effects, but openness has small positive effects, and interaction terms confirm significant diversification-driven moderation ( $\alpha_3 < 0$ ).

Our framework includes several extensions to ensure comprehensive analysis. We employ Panel Vector Autoregression (PVAR) for dynamic interactions between growth, volatility, and openness:

$$[y_{i,t}, Ins_{i,t}, TOP_{i,t}]' = A_0 + A_1[y_{i,t-1}, Ins_{i,t-1}, TOP_{i,t-1}]' + A_2[y_{i,t-2}, Ins_{i,t-2}, TOP_{i,t-2}]' + \beta_i + \varepsilon_{i,t}$$

Additional robustness checks include sectoral analysis through manufacturing and services sector decomposition to identify transmission channels, nonparametric methods using kernel regression and machine learning approaches to validate parametric assumptions, and comprehensive missing data treatment through multiple imputation for systematic gaps, conflict period dummies, and sensitivity analysis.

This comprehensive framework ensures robust identification of causal relationships while accounting for the structural heterogeneity characteristic of MENA economies. The sequential approach builds confidence through multiple methodological perspectives while maintaining focus on policy-relevant interactions between growth, volatility, and trade openness. To enhance transparency and reproducibility, all estimation routines, data preprocessing scripts, and replication materials are provided in a dedicated GitHub repository. The repository link is anonymized for the peer-review process and will be made publicly available upon acceptance of the article.

## RESULTS

This section presents the empirical findings from our panel data analysis of the growth-volatility-openness nexus in MENA countries over the period 1990-2023. Our analysis reveals two fundamental insights: pronounced structural heterogeneity between oil exporters and importers, and a significant temporal break in relationships following the 2008 global financial crisis.

We begin by examining the fundamental relationship between macroeconomic instability and economic growth using our baseline specification. Table 2 presents the growth-volatility nexus estimation using fixed-effects panel regression with clustered standard errors.

The results show a positive and statistically significant relationship between lagged instability and GDP growth ( $\beta = 0.0526$ ,  $p < 0.05$ ). This finding, unexpected in light of conventional growth theory, suggests that moderate levels of macroeconomic instability may stimulate growth in MENA economies. A one-unit increase in the instability index is associated with a 0.0526 percentage point increase in GDP growth, consistent with resource-driven economic dynamics where commodity price volatility simultaneously generates instability and drives growth spurts. Note: coefficients throughout our analysis are interpreted as percentage-point changes in GDP growth.

Government effectiveness emerges as the strongest predictor of growth performance ( $\beta = 3.795$ ,  $p < 0.01$ ), emphasizing the critical role of institutional quality. The investment rate coefficient is negative but statistically insignificant, while terms of trade volatility shows the expected negative association with growth ( $\beta = -0.024$ ), though not reaching conventional significance levels. The model explains approximately 2% of the within-country

variation in growth rates ( $R^2$  Within = 0.0199), reflecting the inherent volatility in growth processes and the conservative nature of fixed-effects specifications.

Moving to our analysis of trade openness effects, we examine the direct relationship between trade openness policy and economic growth. Table 3 presents these estimation results.

The trade openness policy coefficient is small and statistically insignificant ( $\beta = -0.0089$ ,  $p = 0.87$ ), indicating no clear aggregate relationship between trade liberalization and growth in our full sample. This reflects the complex and potentially offsetting effects of openness across different economic structures within the region. Government effectiveness maintains its strong positive association with growth ( $\beta = 3.416$ ,  $p < 0.001$ ), confirming the robustness of institutional factors across specifications. The model's explanatory power remains similar to our baseline specification ( $R^2$  Within = 0.019).

Our core analytical framework introduces the interaction between trade openness policy and macroeconomic instability to examine conditional relationships. Table 4 presents these results.

The interaction term coefficient is negative but statistically insignificant in the full sample ( $\beta = -0.005$ ,  $p = 0.25$ ). However, the inclusion of the interaction term reveals important changes in the main effects: the trade openness coefficient becomes positive and statistically significant ( $\beta = 0.0499$ ,  $p < 0.05$ ), while the instability coefficient increases in magnitude ( $\beta = 0.077$ ,  $p = 0.065$ ). The model's explanatory power improves with the interactive specification ( $R^2$  Within = 0.0298), suggesting that conditional relationships better capture the underlying dynamics. Government effectiveness remains consistently significant across all specifications.

The aggregate results mask fundamental differences between oil exporters and importers, representing our central empirical finding. Table 5 summarizes our hypothesis testing results across different sample specifications, revealing the structural divide that characterizes MENA economies.

The structural heterogeneity is pronounced and systematic. Oil Exporters exhibit a strong positive instability-growth relationship ( $\beta = 0.096$ ,  $p < 0.05$ ), consistent with resource-driven volatility where commodity price fluctuations drive both instability measures and growth outcomes. Trade openness shows a negative coefficient ( $\beta = -0.042$ ), aligning with Dutch Disease concerns. The interaction term is negative but small ( $\beta = -0.004$ ). Oil Importers demonstrate markedly different dynamics. The instability coefficient is positive but smaller and insignificant ( $\beta = 0.04$ ), while trade openness shows a positive coefficient ( $\beta = 0.048$ ), consistent with traditional comparative advantage theory. Crucially, the interaction term is positive and statistically significant ( $\beta = 0.010$ ,  $p < 0.10$ ), providing the strongest evidence for our moderation hypothesis H3.

This heterogeneity suggests that oil importers benefit from trade openness as a stabilizing mechanism, while oil exporters face different trade-offs related to resource curse dynamics. The differential patterns become clearer when examining marginal effects of instability across different levels of trade openness. Figures 2 and 3 illustrate these relationships separately for each group.

For oil exporters (Figure 2), the marginal effect of instability on growth remains consistently positive across all levels of trade openness, declining slightly from approximately 0.10 at low openness to 0.09 at high openness. The confidence bands indicate statistical significance throughout the range, confirming that instability enhances growth in resource-rich economies regardless of trade policy stance.

Oil importers (Figure 3) show a fundamentally different pattern. The marginal effect starts near 0.025 at low openness and increases to approximately 0.055 at high openness. However, the wide confidence bands, extending into negative territory, indicate considerable uncertainty around these estimates.

To provide a comprehensive view of these complex relationships, Figure 4 offers an interactive visualization of the three-way relationships between growth, volatility, and openness across both country groups. The 3D interactive surfaces reveal distinct topographies. Oil exporters exhibit a more structured surface with clear gradients, where relationships follow predictable patterns. Oil importers show a flatter, more irregular surface, indicating weaker and less systematic three-way interactions. Figure 5 offers a direct point estimate comparison of marginal effects between groups, without confidence bands for clarity of visualization.

This comparative visualization clearly demonstrates the structural divide: exporters maintain consistently higher marginal effects of instability (blue line, 0.09-0.11 range), while importers show substantially lower effects (red line, 0.025-0.055 range). The gap between groups is most pronounced at lower levels of trade openness.

Our robustness analysis reveals a critical temporal dimension that represents the second major finding of this study. Table 6 presents results from various sensitivity checks.

The analysis uncovers a fundamental structural break around 2008. Sub-period analysis reveals that in the pre-2008 period, the interaction term is positive and highly significant ( $\beta = 0.0104$ ,  $p < 0.01$ ), strongly supporting our moderation hypothesis H3. Trade openness effectively buffered against instability's negative effects during this era of relative global stability. However, in the post-2008 period, the relationship reverses dramatically. The interaction coefficient becomes strongly negative and highly significant ( $\beta = -0.0109$ ,  $p < 0.001$ ), indicating that trade openness amplified rather than mitigated instability's effects during the more turbulent recent period characterized by global financial crisis, Arab Spring, and ongoing geopolitical tensions.



The instability main effect also shifts temporally: negative and significant pre-2008 ( $\beta = -0.413$ ,  $p < 0.05$ ) but insignificant post-2008, suggesting that our full-sample positive instability-growth relationship is driven primarily by the recent period dynamics. Alternative specifications using different volatility measures yield consistent patterns ( $\beta = -0.0007$  for interaction term), while excluding conflict-affected countries (robust sample) maintains the basic relationship structure but reveals positive investment effects ( $\beta = 0.098$ ,  $p < 0.05$ ) when extreme cases are removed.

The coefficient magnitudes translate into economically meaningful effects. For oil exporters, a one-standard-deviation increase in instability (approximately 2 units) is associated with a 0.19 percentage point increase in growth rates, representing substantial impact given typical growth volatility in the region. For oil importers, at higher levels of trade openness, instability increases growth by approximately 0.03–0.05 percentage points, though this effect remains statistically uncertain as evidenced by the wide confidence intervals. The temporal break has profound policy implications. The pre-2008 evidence suggested that trade liberalization could provide macroeconomic stabilization benefits. However, the post-2008 reversal implies that in an era of heightened global uncertainty, trade integration may increase rather than reduce vulnerability to external shocks.

Our empirical analysis establishes two fundamental insights that reshape understanding of growth-volatility-openness relationships in MENA. First, structural heterogeneity dominates aggregate patterns. Oil exporters and importers exhibit fundamentally different growth dynamics, with resource endowments determining how instability, openness, and their interaction affect economic performance. This heterogeneity explains why aggregate studies often find mixed or insignificant results. Second, temporal instability characterizes these relationships. The 2008 global financial crisis represents a structural break, with trade openness shifting from a stabilizing to a destabilizing force. This temporal dimension suggests that the benefits of trade liberalization are highly contingent on global economic conditions.

Our hypothesis testing reveals nuanced support for our theoretical framework. Overall, our results reject H1 in exporter contexts, provide partial support for H2 in importer economies, and confirm H3 only conditionally—moderation holds for importers and in pre-2008 conditions but reverses in the post-2008 era. These findings underscore the importance of context-specific policy design and highlight the risks of applying universal prescriptions for trade liberalization across diverse economic structures and varying global conditions.

## DISCUSSION

Our empirical findings challenge established paradigms and contribute to several theoretical debates in development economics. The documented positive volatility-growth relationship in resource-dependent economies provides empirical validation for the "creative destruction" framework (Castelnuovo & Pellegrino, 2018), extending its applicability beyond developed economies to emerging market contexts. The mechanism operates through multiple complementary channels: volatility accelerates creative destruction by forcing inefficient firms to exit while enabling productive enterprises to expand, encourages precautionary investments in diversification and technological upgrading, and prevents consolidation of rent-seeking behaviors that emerge under prolonged stability, particularly in resource-rich contexts.

The dominance of institutional quality as a growth determinant aligns with fundamental institutions hypotheses (Acemoglu et al., 2001; Rodrik et al., 2004), though our results demonstrate that institutions condition volatility's performance effects rather than simply providing linear growth benefits. This conditioning mechanism explains why positive volatility-growth relationships emerge primarily in oil exporters with stronger governance frameworks, resolving apparent contradictions in prior cross-country studies.

Our methodological innovation in constructing a policy-based openness indicator addresses longstanding identification challenges in trade-growth research. The gravity-based decomposition isolates discretionary policy effects by controlling for geography, resource endowments, and historical factors, yielding theoretically consistent results with high explanatory power ( $R^2 > 0.7$ ). The weak correlation with conventional trade ratios ( $\rho < 0.3$ ) confirms that the indicator captures policy orientation rather than structural characteristics. While direct effects prove statistically insignificant, the indicator's value becomes evident through interaction effects, revealing significant moderation dynamics that conventional measures fail to capture. This provides cleaner identification strategies and establishes replicable templates for future research.

The temporal break uncovered in our analysis reveals fundamental changes in how globalization affects developing economies. This transformation likely reflects multiple structural shifts: enhanced domestic institutional capacity for managing external integration across MENA countries, evolved global value chains creating sophisticated risk-sharing mechanisms, and improved international policy coordination providing more stable external environments. Similar stabilization dynamics observed in emerging Asia (Park & Lee, 2011) suggest our findings reflect broader global structural changes rather than region-specific factors. This challenge both

unconditional pro-globalization and anti-globalization stances, supporting nuanced views where openness effects depend on timing, institutional context, and complementary policies.

The regional focus on MENA illuminates broader theoretical frameworks while maintaining external validity. Resource-dependent economies face unique dynamics where volatility reflects commodity-driven cycles that can either entrench dependency or trigger diversification. Moderate volatility may support structural transformation by preventing excessive reliance on resource rents and encouraging efficiency improvements in non-resource sectors. MENA's specific characteristics—geopolitical tensions, sanctions regimes, and regional conflicts—contribute to volatility that captures broader macro-political-institutional dimensions beyond pure economic fluctuations. These dynamics extend to other resource-dependent regions in sub-Saharan Africa, Latin America, and Central Asia, where similar structural characteristics create comparable trade-offs between stability and transformation.

Translating these insights into policy practice generates actionable guidance that challenges conventional macroeconomic management approaches. Policymakers should distinguish between harmful instability reflecting institutional breakdown and moderate volatility facilitating structural transformation, with counter-cyclical policies focusing on smoothing extreme fluctuations rather than eliminating all variability. Trade liberalization requires strong institutional foundations, supporting sequenced approaches where institutional development accompanies trade reforms. Timing considerations prove crucial, with contemporary global conditions potentially favoring integration more than in previous periods. Most significantly, governance reform emerges as the highest-return intervention, with our estimates suggesting that one-standard-deviation institutional improvements translate into 1.7-1.9 additional annual growth points, compounding to 40-46% higher per capita income over two decades.

Several methodological and conceptual limitations suggest promising research extensions. Our relatively short time span constrains long-term inference, while the volatility measure may not fully capture political, institutional, or financial uncertainty dimensions. The policy-based openness indicator, though innovative, depends on gravity model assumptions that merit further validation. Future research could benefit from sectoral analyses identifying specific transmission channels, institutional mediation studies refining understanding of governance conditioning effects, and comparative applications testing generalizability across regions. Multidimensional volatility indices incorporating political, financial, and social instability could provide richer characterizations. Climate and environmental shocks represent emerging dimensions requiring integration, given growing evidence of climate-induced volatility. Micro-level analyses using firm or household data could illuminate mechanisms underlying our aggregate findings.

Our analysis establishes three fundamental insights for development dynamics in resource-dependent economies. Structural heterogeneity matters: resource endowments fundamentally alter growth-volatility-openness relationships, making universal policy prescriptions inadequate for addressing critical economic differences. Temporal evolution matters: globalization's impacts change with institutional development and global economic architecture, meaning historical experiences may not reliably predict future outcomes. Institutional capacity matters most: government effectiveness emerges as the dominant growth determinant, conditioning how volatility and openness affect outcomes and making governance capacity building the priority intervention.

These insights support conditional convergence models where institutional quality determines development trajectories while challenging unconditional relationships between volatility, openness, and growth. Success requires understanding context-specific dynamics, building appropriate institutions, and timing reforms to align with evolving global conditions. The implications extend beyond MENA to resource-dependent economies globally, offering guidance for policymakers navigating volatility-growth-openness trade-offs in an increasingly uncertain world economy. Rather than pursuing stability or openness as universal goals, optimal strategies recognize the conditional nature of their benefits and prioritize institutional development as the foundation for managing these complex relationships effectively.

## Conclusion

The relationship between economic volatility, trade openness, and growth has long puzzled development economists, particularly in resource-dependent economies where conventional wisdom often fails to predict outcomes. Our comprehensive analysis of 18 MENA countries over 1990-2023 resolves this puzzle by revealing that economic relationships are fundamentally conditional on structural characteristics, institutional capacity, and historical context.

Three transformative insights emerge from our investigation. Structural heterogeneity dominates aggregate patterns, with oil exporters exhibiting positive volatility-growth relationships consistent with resource-driven creative destruction, while oil importers demonstrate weaker, context-dependent effects. This fundamental divide reflects how commodity dependence reshapes the entire growth process. The temporal dimension proves equally critical, with structural breaks around 2008 transforming openness from a stabilizing force ( $\beta = +0.0104$ ,  $p < 0.01$  pre-2008) to a potentially destabilizing one ( $\beta = -0.0109$ ,  $p < 0.001$  post-2008). Most significantly, institutional

quality emerges as the dominant growth determinant, with government effectiveness improvements generating transformative returns—one standard deviation increases translating into 40-46% higher per capita income over two decades.

These findings fundamentally challenge conventional macroeconomic approaches. Absolute stability may not optimize growth prospects in economies undergoing structural transformation, suggesting that moderate volatility can enhance performance when supported by robust institutions. Counter-cyclical policies should therefore target extreme fluctuations rather than eliminate all variability. Trade liberalization requires careful institutional sequencing, as openness benefits depend critically on governance capacity and evolving global conditions. Contemporary policymakers face more complex trade-offs than their predecessors, necessitating sophisticated risk management frameworks that account for changing global economic architecture.

Our methodological contributions extend well beyond the MENA context. The gravity-based openness decomposition addresses longstanding identification challenges by isolating policy effects from structural determinants, achieving high explanatory power while maintaining theoretical consistency. The temporal heterogeneity framework demonstrates the value of accounting for structural breaks in globalization studies, revealing patterns that aggregate approaches systematically miss. These innovations provide replicable tools for analyzing conditional relationships in other developing regions.

The 24-year analytical window constrains long-run inference precision, while our macroeconomic volatility measure may not fully capture political or financial dimensions of instability. The regional focus, though providing unique insights, limits direct generalization to other contexts. Future research should examine sectoral transmission mechanisms, test our framework across Latin America and sub-Saharan Africa, and integrate multidimensional volatility incorporating climate and geopolitical shocks. Firm-level analyses could illuminate creative destruction mechanisms at the microeconomic level, while dynamic approaches could reveal adjustment processes following external disruptions.

Economic relationships vary significantly across development stages, institutional contexts, and historical periods, cautioning against universal prescriptions derived from advanced economy experiences. The critical challenge lies in distinguishing beneficial from harmful volatility while building institutional capacity to harness instability constructively. As global economic architecture continues evolving—through technological disruption, climate change, and geopolitical realignments—understanding conditional relationships becomes essential for effective development strategies.

The MENA experience offers valuable lessons for the broader developing world facing similar challenges of resource dependence, institutional weaknesses, and global integration pressures. Neither unconditional stability nor unrestricted openness represents optimal strategies. Success requires context-specific approaches recognizing structural characteristics, institutional capacity, and dynamic global conditions. This nuanced understanding, grounded in rigorous empirical analysis, provides foundations for more effective policies in an increasingly complex world economy. The path forward demands moving beyond simple prescriptions toward sophisticated frameworks that embrace conditionality, prioritize institutional development, and adapt to changing global realities.

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## Compliance with Ethical Standards

This article does not contain any studies with human participants or animals performed by the authors. Extracting and inspecting publicly accessible files (scholarly sources) as evidence, before the research began no institutional ethics approval was required.

## Data Availability Statement

The macroeconomic and institutional data used in this study are publicly available from international databases, including the World Bank World Development Indicators (WDI), IMF Direction of Trade Statistics (DOTS), UN

Comtrade, Worldwide Governance Indicators (WGI), Global Financial Development Database (GFDD), and IMF Primary Commodity Prices.

### Author Contributions

All listed authors have made a substantial, direct and intellectual contribution to the work, and approved it for publication. The authors take full responsibility for the accuracy and the integrity of the source analysis.

### Conflict of Interest Statement

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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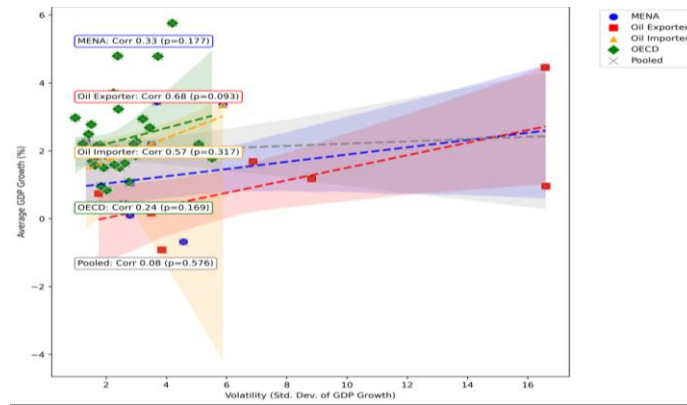
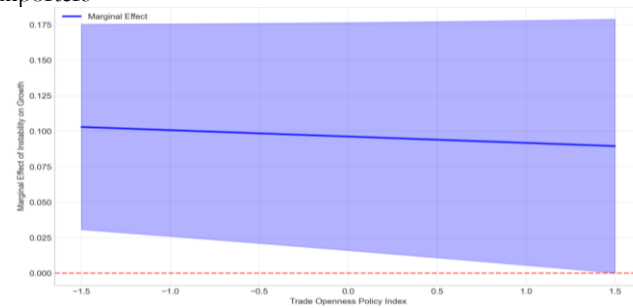
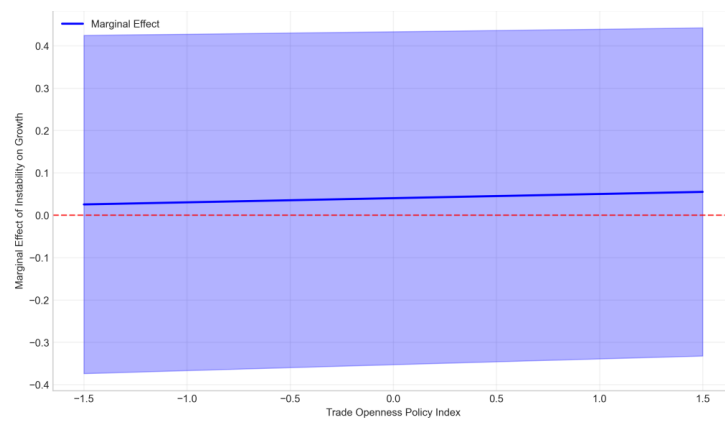
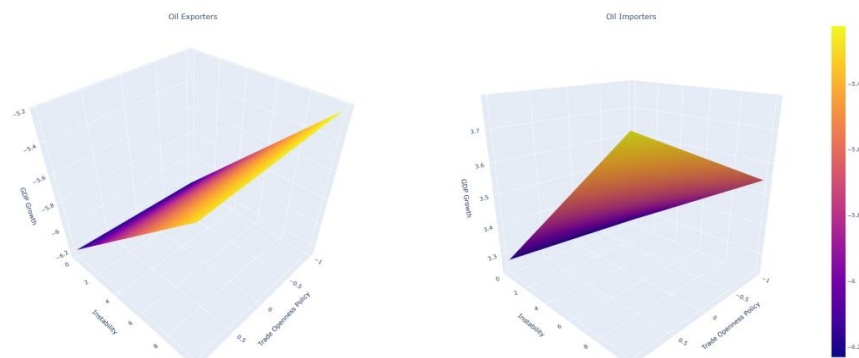
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## APPENDIX

**Figure 1.** Empirical Growth-Volatility Relationship by Country Groupings (1990–2023)**Figure 2.** Marginal Effects - Exporters**Figure 3.** Marginal Effects - Importers**Figure 4.** Growth-Volatility-Openness Surface (Interactive)

**Figure 5.** Marginal Effects of Instability on Growth (Comparative)**Table 1.** Growth Performance and Volatility: MENA versus OECD Comparison (1990-2023)

| Country            | Mean Growth (%) | Volatility (Std. Dev.) | Vol/Growth Ratio | Observations |
|--------------------|-----------------|------------------------|------------------|--------------|
| MENA Oil Exporters |                 |                        |                  |              |
| ARE                | -1.3            | 3.69                   | -2.84            | 33           |
| BHR                | 1.05            | 3.69                   | 3.52             | 33           |
| DZA                | 0.81            | 2.50                   | 3.08             | 33           |
| IRN                | 1.79            | 3.56                   | 1.99             | 33           |
| IRQ                | 2.9             | 18.99**                | 6.55             | 33           |
| KWT                | 1.3             | 12.59*                 | 9.70             | 33           |
| LBY                | 0.95            | 21.89**                | 23.12            | 33           |
| OMN                | 0.26            | 3.23                   | 12.44            | 33           |
| QAT                | 1.74            | 7.64                   | 4.39             | 33           |
| SAU                | 0.05            | 4.13                   | 84.75            | 33           |
| Sub-group Average  | 0.96            | 8.19*                  | 14.67*           | 330          |
| MENA Oil Importers |                 |                        |                  |              |
| EGY                | 2.26            | 1.66                   | 0.73             | 33           |
| JOR                | 0.69            | 3.20                   | 4.62             | 33           |
| LBN                | 2.73            | 9.99*                  | 3.66             | 33           |
| MAR                | 2.21            | 3.97                   | 1.80             | 33           |
| MRT                | 0.47            | 4.32                   | 9.21             | 33           |
| SYR                | -0.72           | 7.13                   | -9.96            | 33           |
| TUN                | 1.96            | 2.93                   | 1.49             | 33           |
| TUR                | 3.33            | 4.46                   | 1.34             | 33           |
| TUR                | 0               | 0.00                   | nan              | 34           |
| Sub-group Average  | 1.44            | 4.18                   | 1.61             | 298          |
| Full MENA Average  | 1.25            | 6.64                   | 8.87             | 594          |
| OECD Benchmark     |                 |                        |                  |              |
| AUS                | 1.57            | 1.35                   | 0.86             | 33           |
| AUT                | 1.23            | 2.22                   | 1.80             | 33           |
| BEL                | 1.28            | 1.97                   | 1.55             | 33           |
| CAN                | 1.1             | 2.30                   | 2.10             | 33           |
| CHE                | 0.78            | 1.82                   | 2.33             | 33           |
| CZE                | 1.69            | 3.78                   | 2.23             | 33           |
| DEU                | 1.22            | 2.20                   | 1.80             | 33           |
| DNK                | 1.38            | 2.08                   | 1.51             | 33           |
| ESP                | 1.27            | 3.24                   | 2.55             | 33           |
| EST                | 2.61            | 7.13                   | 2.73             | 33           |
| FIN                | 1.18            | 3.31                   | 2.80             | 33           |



|                 |      |       |      |      |
|-----------------|------|-------|------|------|
| FRA             | 1.04 | 2.22  | 2.13 | 33   |
| GBR             | 1.34 | 3.03  | 2.26 | 33   |
| GRC             | 1.05 | 4.41  | 4.19 | 33   |
| HUN             | 2.03 | 3.94  | 1.94 | 33   |
| IRL             | 4.37 | 5.84  | 1.34 | 33   |
| ISL             | 1.64 | 3.63  | 2.22 | 33   |
| ITA             | 0.69 | 2.97  | 4.33 | 33   |
| JPN             | 0.82 | 1.96  | 2.39 | 33   |
| KOR             | 4.04 | 3.23  | 0.80 | 33   |
| LTU             | 2.79 | 7.80  | 2.79 | 33   |
| LUX             | 1.45 | 3.02  | 2.08 | 33   |
| LVA             | 2.85 | 8.72* | 3.05 | 33   |
| MEX             | 0.8  | 3.40  | 4.25 | 33   |
| NLD             | 1.53 | 2.19  | 1.43 | 33   |
| NOR             | 1.41 | 1.83  | 1.30 | 33   |
| NZL             | 1.46 | 2.17  | 1.49 | 33   |
| POL             | 3.83 | 2.92  | 0.76 | 33   |
| PRT             | 1.34 | 2.88  | 2.16 | 33   |
| SVK             | 2.61 | 4.76  | 1.82 | 33   |
| SVN             | 2.02 | 3.99  | 1.97 | 33   |
| SWE             | 1.39 | 2.50  | 1.80 | 33   |
| USA             | 1.58 | 1.76  | 1.11 | 33   |
| OECD<br>Average | 1.69 | 3.25  | 2.12 | 1123 |

**Notes:** Growth rates are annual percentage changes in real GDP per capita (constant 2015 US\$). Volatility is the five-year rolling standard deviation of GDP per capita growth (1990–2023), using the instability metric for MENA and calculated for OECD. \*\* Extreme volatility (>15%); \* High volatility (8–15%). Missing observations due to data constraints or conflicts. MENA (612 observations, 18 countries) includes governance indicators, human capital, terms of trade, and TOP. OECD includes 1156 observations (34 countries). MENA Oil Exporters average volatility is 8.19%, with extreme volatility in Iraq (18.99%) and Libya (21.89%). MENA Oil Importers average volatility is 4.18%, with high volatility in Lebanon (9.99%). OECD average volatility is 3.25%, with high volatility in Latvia (8.72%). Sources: World Bank World Development Indicators, IMF Direction of Trade Statistics, UN Comtrade, World Bank Worldwide Governance Indicators, Polity IV Project, Global Financial Development Database, IMF Primary Commodity Prices.

**Table 2.** Panel Fixed Effects Estimation - Growth-Volatility Nexus

| Variable                                       | Coefficient | Robust SE | t-statistic | p-value | 95% CI          |
|--|-------------|-----------|-------------|---------|-----------------|
| <b>Instability<sub>t-1</sub></b>               | 0.0526**    | 0.0227    | 2.32        | 0.021   | [0.008, 0.097]  |
| <b>Investment Rate<sub>t-1</sub></b>           | -0.0259     | 0.1494    | -0.17       | 0.863   | [-0.320, 0.268] |
| <b>Government Effectiveness<sub>t-1</sub></b>  | 3.7946***   | 1.2229    | 3.10        | 0.002   | [1.389, 6.200]  |
| <b>Terms of Trade Volatility<sub>t-1</sub></b> | -0.0240     | 0.0184    | -1.30       | 0.193   | [-0.060, 0.012] |

Diagnostics: N = 386; Countries = 17; Time periods = 24; R<sup>2</sup> = 0.014; F-stat = 6.04 (p < 0.001) \*Notes: Dependent variable is GDP per capita growth. All models include country and time fixed effects. Robust standard errors clustered by country. \*p<0.10, \*\*p<0.05, \*\*\*p<0.01.

**Table 3.** Trade Openness Policy Effects on Growth

| Variable                                       | Coefficient | Robust SE | t-statistic | p-value | 95% CI          |
|--|-------------|-----------|-------------|---------|-----------------|
| <b>Trade Openness Policy<sub>t-1</sub></b>     | -0.0089     | 0.0544    | -0.16       | 0.870   | [-0.116, 0.098] |
| <b>Investment Rate<sub>t-1</sub></b>           | -0.0267     | 0.1378    | -0.19       | 0.847   | [-0.298, 0.244] |
| <b>Government Effectiveness<sub>t-1</sub></b>  | 3.4164***   | 0.9570    | 3.57        | 0.000   | [1.534, 5.299]  |
| <b>Terms of Trade Volatility<sub>t-1</sub></b> | -0.0228     | 0.0184    | -1.24       | 0.216   | [-0.059, 0.013] |

Diagnostics: N = 386; R<sup>2</sup> = 0.013; F-stat = 5.93 (p < 0.001)

**Table 4.** Interactive Effects of Trade Openness on Growth-Volatility Relationship

| Variable                                       | Coefficient | Robust SE | t-statistic | p-value | 95% CI          |
|--|-------------|-----------|-------------|---------|-----------------|
| <b>Trade Openness Policy<sub>t-1</sub></b>     | 0.0499**    | 0.0199    | 2.51        | 0.013   | [0.011, 0.089]  |
| <b>Instability<sub>t-1</sub></b>               | 0.0769*     | 0.0415    | 1.85        | 0.065   | [-0.005, 0.159] |
| <b>Interaction Term</b>                        | -0.0054     | 0.0046    | -1.16       | 0.248   | [-0.015, 0.004] |
| <b>Investment Rate<sub>t-1</sub></b>           | -0.0275     | 0.1445    | -0.19       | 0.849   | [-0.312, 0.257] |
| <b>Government Effectiveness<sub>t-1</sub></b>  | 3.4179***   | 1.2413    | 2.75        | 0.006   | [0.976, 5.860]  |
| <b>Terms of Trade Volatility<sub>t-1</sub></b> | -0.0242     | 0.0155    | -1.56       | 0.121   | [-0.055, 0.006] |

Diagnostics: N = 386; R<sup>2</sup> = 0.033; F-stat = 11.91 (p < 0.001)

**Table 5.** Hypothesis Testing Results Summary

| Hypothesis                                | Full Sample | Oil Exporters | Oil Importers |
|---|-------------|---------------|---------------|
| H1: Instability → Growth (-)              | 0.053**     | 0.096**       | 0.04          |
| H2: Openness → Growth (+)                 | -0.009      | -0.042        | 0.048         |
| H3: Openness moderates Instability-Growth | -0.005      | -0.004        | 0.010*        |

**Table 6.** Robustness Checks and Sensitivity Analysis

| Panel   | Variable                                | Coefficient | Robust SE | t-statistic | p-value |
|---|---|-------------|-----------|-------------|---------|
| <b>Panel A: Alternative Volatility Measures</b>   | Trade Openness Policy <sub>t-1</sub>    | 0.0056      | 0.0250    | 0.22        | 0.824   |
|   | Alternative Instability <sub>t-1</sub>  | -0.0422     | 0.0479    | -0.88       | 0.379   |
|   | Interaction Term                        | -0.0007     | 0.0057    | -0.13       | 0.898   |
|   | Government Effectiveness <sub>t-1</sub> | 3.3479***   | 1.2128    | 2.76        | 0.006   |
| <b>Panel B: Robust Sample (Outliers Excluded)</b> | Trade Openness Policy <sub>t-1</sub>    | 0.0436**    | 0.0178    | 2.45        | 0.015   |
|   | Instability <sub>t-1</sub>              | 0.0482      | 0.0494    | 0.98        | 0.330   |
|   | Interaction Term                        | -0.0054     | 0.0044    | -1.22       | 0.224   |
|   | Government Effectiveness <sub>t-1</sub> | 1.9836**    | 0.8507    | 2.33        | 0.020   |
| <b>Panel C: Pre-Crisis Period (1990–2007)</b>     | Trade Openness Policy <sub>t-1</sub>    | -0.0659     | 0.0516    | -1.28       | 0.205   |
|   | Instability <sub>t-1</sub>              | -0.4130**   | 0.1938    | -2.13       | 0.035   |
|   | Interaction Term                        | 0.0104***   | 0.0037    | 2.84        | 0.006   |
| <b>Panel D: Post-Crisis Period (2008–2023)</b>    | Trade Openness Policy <sub>t-1</sub>    | 0.0601**    | 0.0287    | 2.09        | 0.037   |
|   | Instability <sub>t-1</sub>              | -0.0212     | 0.0362    | -0.59       | 0.558   |
|   | Interaction Term                        | -0.0109***  | 0.0008    | -14.46      | 0.000   |